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# Roe deer population structure in a highly fragmented landscape

Peter Breyne · Joachim Mergeay · Jim Casaer

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**Abstract** Northern Belgium (Flanders) is one of the most densely populated and urbanized regions in Europe. Many species are therefore likely to suffer from anthropogenic pressure and habitat destruction and fragmentation. Although many large mammals are recolonizing in parts of Europe, including Belgium, due to adaptation, a relaxation of persecution and habitat restoration, we have little actual data concerning the effects of landscape features on their population structure. We analysed the genetic structure of discrete roe deer (*Capreolus capreolus*) populations in the Eastern part of Flanders, with special emphasis on the impact of habitat fragmentation and anthropogenic barriers. The sampled populations were clearly genetically differentiated. Genetic structure could be explained by purely distance-based landscape modelling, but a simpler model focusing solely on barrier effects of large transportation infrastructure explained nearly as much genetic variance. In contrast, analyses based on least-cost landscape modelling failed to yield a significant effect. Overall, the results suggest considerable landscape-level effects of transportation infrastructure.

**Keywords** *Capreolus capreolus* · Fragmentation · Genetic drift · Redundancy analysis · Landscape genetics

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## Introduction

The Western European landscape has changed drastically during the last decades resulting in both landscape homogenization and fragmentation. The increasing road density and intensity of its use, and the building of new railroads, have led to an increase of barriers in the landscape (Antrop 2004). Several mammals are hampered in their movements by urban areas, roads and unattractive land (Jongman 2002; Frantz et al. 2012), while habitat destruction due to anthropogenic pressures results in both habitat loss and habitat fragmentation (Bright 1993; Fahrig 2003). Though habitat fragmentation per se can have both positive and negative effects on animal abundance and surviving, the net outcome is usually negative (Andren 1994, Jaeger et al. 2011). In Flanders, the northern part of Belgium, anthropogenic pressures have resulted in one of the highest habitat fragmentation levels of Europe (Jaeger et al. 2011). Flanders is one of the most densely populated regions worldwide (462 inhabitants km<sup>-2</sup>) with scattered urban areas (25 % of the surface) alternating with agricultural areas (62 % of the surface). Forested areas represent around 10 % of the total surface. With a road density of around 5 km km<sup>2</sup> and many railways, linear fragmentation of the landscape is extremely significant (Tromc e et al. 2002). Furthermore, the landscape is intersected by a series of canals with steep borders, making them important barriers even for larger mammals capable of swimming. A landscape fragmentation study in Europe (Jaeger et al. 2011) showed that Northern Belgium, together with the adjacent parts of north-western France (French Flanders) and western Germany (North Rhine Westphalia) can be considered as a single uninterrupted stretch of extremely fragmented landscape. For many wildlife species in this area, this likely result in small isolated populations where dispersal, necessary for gene flow and (re)colonisation after local extinction, is strongly impeded (Hanski 1998).

While the impact of land use and landscape features on animal dispersal and gene flow is widely accepted, it is only with the emergence of studies within the field of landscape genetics that a direct link between landscape features and genetic structure could be quantified and proven (Manel et al. 2003; Holderegger and Wagner 2008). Several studies have shown that anthropogenic barriers and unsuitable landscape features impede gene flow (Fahrig 2003; Epps et al. 2005; Balkenhol and Waits 2009; Frantz et al. 2012) even for medium-sized ungulate species such as roe deer which are expected to be very mobile (Coulon et al. 2006; Kuehn et al. 2007). In spite of the heavy anthropogenic pressure, the high mobility and dispersal capacity of several medium-sized mammals are exemplified by the recent surge of recolonization of previously abandoned regions in Western Europe. Wolves (*Canis lupus*), Eurasian lynxes (*Lynx lynx*), Eurasian otters (*Lutra lutra*) and more common species such as wild boars (*Sus scrofa*) and beavers (*Castor fiber*) have re-expanded their ranges in the last decades, some of them even into Flanders (Deinet et al. 2013).

Roe deer is currently the most abundant native free-living ungulate species in Flanders, although this is a recent phenomenon. Roe deer densities and distribution strongly increased after the Second World War. Its distribution in the 1960s was still limited in the northern part of the provinces of Limburg and Antwerp while some relict populations were found in the central part of the country around Brussels and Leuven (Casaer and Licoppe 2010). Most probably, roe deer never totally disappeared from most of these regions. In the 1980s, the species started to recover from its earlier decline, with gradual recolonization of previously abandoned areas (Casaer 2003). Currently, it has a broader distribution range than ever registered before, with local population densities exceeding 30 animals km<sup>-2</sup> of forested area (Casaer and Licoppe 2010). The expansion can be mainly explained by the lack of predators, changes in the hunting regulation, abundant food availability (year round access to arable land) and the high adaptation capacity of roe deer to new habitat conditions. Undocumented translocations and introduction of animals may also have contributed to the current distribution. Roe deer primarily inhabits woodland areas although it occupies also more open landscapes with woody structures such as hedgerows (Morellet et al. 2011). In fragmented landscapes, roe deer stick to wooded patches and show high site fidelity (Hewison et al. 2001; Cargnelutti et al. 2002). Animals live solitary or in small family groups and exhibit a low level of polygyny. Both sub adult males and females disperse in spring or early summer of the year following birth or in the subsequent spring (Wahlström and Liberg 1995; Linnell et al. 1998). Typical dispersal distances are thought to be low and are mostly restricted to a few kilometres (Linnell et al. 1998), although little is known about dispersal in highly fragmented landscapes. Such a pattern of small-scale dispersal should lead

to increasing genetic differentiation with geographic distance (isolation-by-distance). However, in highly fragmented landscapes, it was found that for roe deer genetic distances correlate with urbanization rather than with geographic distance (Wang and Schreiber 2001). The relatively rapid recolonization of formerly abandoned regions nevertheless indicates that the potential to disperse is high, but gene flow dynamics among established populations are potentially low due to behavioural and life history constraints (e.g., Vangestel et al. 2011).

Here, we attempt to identify the important determinants and causes of genetic structuring of roe deer populations across the study region. First, we describe the population genetic structure of roe deer populations in Flanders. We then test which landscape description fits the genetic data best: purely distance-based spatial modelling, the position of populations relative to major transportation infrastructure, or a least-cost analysis that integrates the effect of green landscape features.

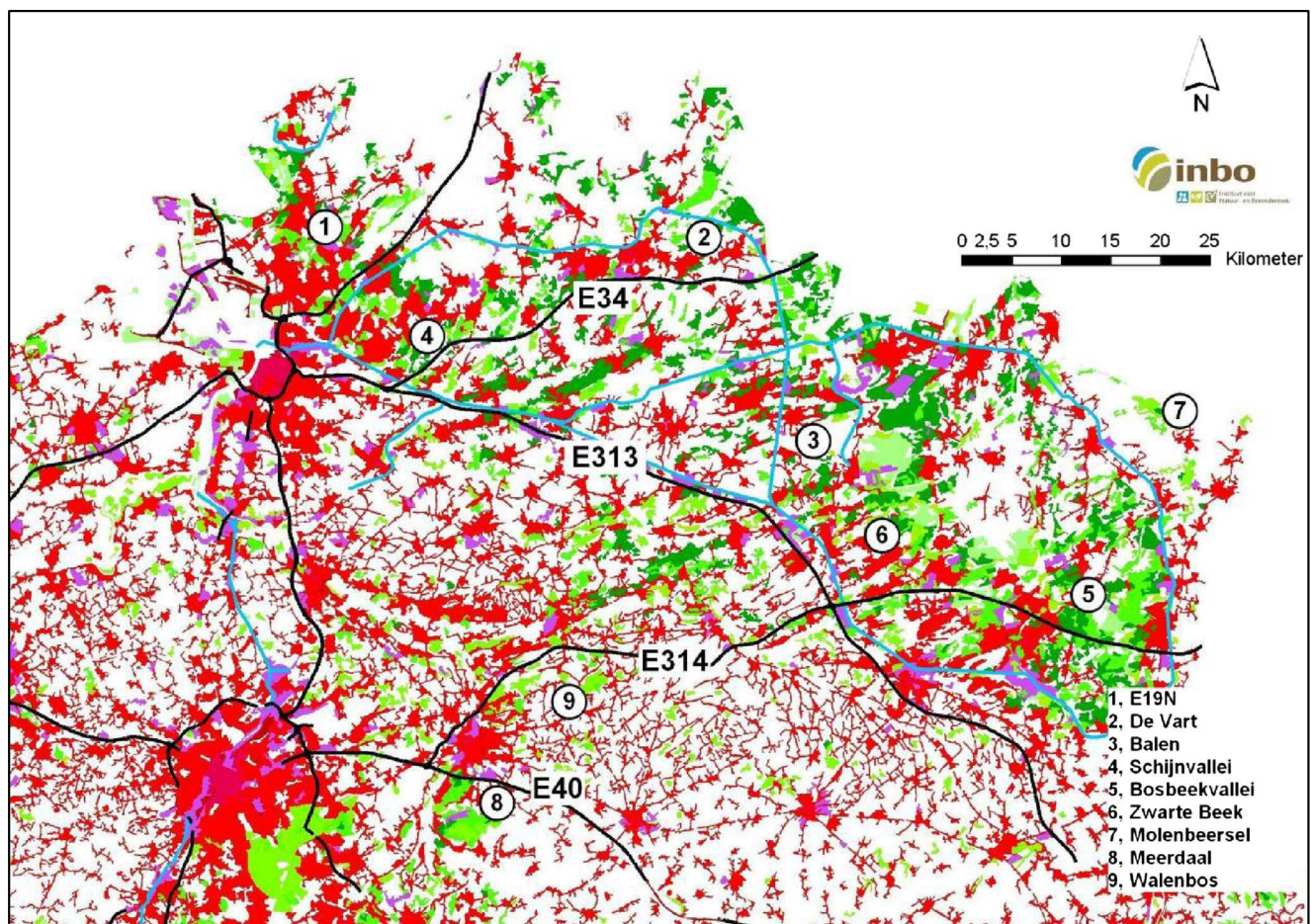
## Methods

### Study area and sampling

The study area is situated in the eastern part of Flanders (Fig. 1) and covers approximately 14,000 km<sup>2</sup> of which 20 % is forested. The study area is intersected by several highways, railways and canals and contains many cities and urbanized regions (Fig. 1). Samples were obtained from nine Game Management Units (GMUs) in the Flemish region (Fig. 1 and Online Resource 1), ranging in size between 19.5 and 99.8 km<sup>2</sup> with a mean of 55.9 km<sup>2</sup>. One GMU (Meerdaal) is intersected by a highway (see Online Resource 1), but only animals from the south of the highway were used in this study. As a reference for a stable and large population, samples were included from the region of Elsenborn, located in the Walloon region, about 70 km away from the nearest Flemish sample location. Muscle tissue samples from legally shot animals were collected between 2005 and 2007, and were preserved in 95 % ethanol. More detailed information on sex and age of the samples is given in Online Resource 2.

### Genetic analysis

DNA extraction was performed with the DNeasy Tissue Kit (Qiagen) according to the manufacturer's protocol. Ten microsatellite loci (BM757, OarFCB304, BMC1009, HUI1177, NVHRT48, CSSM41, CSSM43, IDVGA29, CSSM39 and BM1706) were amplified in a single multiplex reaction as described by Galan et al. (2003). PCR amplifications were analysed by capillary electrophoresis on a SCE 9610 genetic



**Fig. 1** Sampling area in Eastern Flanders. Urbanized regions are coloured in *red*, industrial zones in *purple*, highways in *black* and canals in *blue*. The names of the highways mentioned in the text are indicated.

Forests and other suitable habitats (marshland, heathland,...) are shown in different shades of green

analyzer (Spectrumedix) using the Genospectrum 3.0.0 software.

#### Data analysis

##### *Genetic diversity and population structure*

FSTAT v.2.9.3 (Goudet 2001) was used to infer the number of alleles (NA) and allelic richness (AR) per locus and population as well as to calculate  $F_{is}$  and  $F_{st}$  values (900 permutations). As alternative for measuring differentiation,  $D_{est}$ -values (Jost 2008) were calculated using the programme SMOGD (Crawford 2010) with 999 bootstraps. Observed ( $H_o$ ) and expected ( $H_e$ ) heterozygosity were determined according to Nei (1978) using GENETIX v.4.03 (Belkhir 2004). As test for deviation from Hardy-Weinberg equilibrium (HWE), significance of heterozygote deficiency was calculated using GENEPOP v.4.0.10 (Rousset 2008) with 10,000 dememorizations, 300 batches and 10,000 iterations per batch. To visualize potential population structure, a principal coordinate analysis (PCoA) was performed in GENALEX v.6.41 (Peakall and Smouse 2006). The Bayesian

assignment approach implemented in STRUCTURE 2.3.2 (Pritchard et al. 2000; Falush et al. 2003) was used to infer the number of genetic clusters in the total sample. For each of  $K=1$  to 11, a model assuming admixture and correlated allele frequencies was run five times with a burn-in period of 100,000 followed by 400,000 MCMC iterations, with prior spatial information provided except for the first run. The genetic structure was inferred based on the mean  $\ln P(D)$  and by calculating  $\Delta K$  (Evanno et al. 2005) using STRUCTURE HARVESTER (Earl and von Holdt 2011). GENECLASS2 (Piry et al. 2004) was used to assign individuals to predefined populations (self-assignment) following the partial Bayesian approach of Rannala and Mountain (1997). The Monte Carlo method of Paetkau et al. (2004) was used for exclusion analysis applying default settings.

##### *Spatial genetic structure*

We used redundancy analysis (RDA), a constrained ordination method (see Legendre and Legendre 2012 for a comprehensive overview), to analyze the genetic structure of populations (dependent variables) relative to three explanatory variables: (1)

the spatial position in the landscape, (2) major transportation infrastructure in the region and (3) three measures of least-cost distances. The first model adopts distance as the primary driver of genetic structure (isolation by distance), the second assumes that highways and canals are the primary drivers of genetic structure (isolation by barriers) and the third that both distance and the permeability of different landscape features matter.

Least cost path analysis was performed in ARCGIS v.10 (ESRI) using the Spatial Analyst tools. A cost raster was first created using a land use map (Gobin et al. 2009) with cells of  $15 \times 15$  m and assigning costs to each cell based on the presence of predefined barriers and landscape features (Table 1). We preferred to work with such small cell sizes as it is easier to deal with linear barrier elements such as high ways and canals, what was very important in our study given the fact that the study area contains many barriers. To weight the effect of geographical distances to a varying degree when calculating the least cost path, three least cost matrices (LCM) were generated with cost values of 1 to 10 (1, no resistance; 5, intermediate resistance and 10, absolute barrier; LCM10), 1 to 100 (1, no resistance; 50, intermediate resistance and 100, absolute barrier; LCM100) and 1 to 1000 (1, no resistance; 500, intermediate resistance and 1000, absolute barrier; LCM1000). The cost value of each landscape feature was determined based on the estimated permeability or barrier effect for roe deer (Table 1). Next, a cost distance raster was created for each of the sampling areas (polygon) equalling per raster cell the cumulative cost of reaching the centre of the polygon starting from that cell. As we had no individual coordinates of the shot animals, we used the polygon centroids of each sampling area as geographical coordinates. This might have caused some imprecision, but since most distances between the different game management units are far much bigger than the distances between individual animals within a unit, we didn't expect too much bias. Finally, least cost paths between each pair of sampling areas were calculated, resulting in the final least cost path matrices.

**Table 1** Predefined barriers and landscape features and assigned cost values

Barrier/landscape feature	Cost value
Residential and commercial area	10, 100, 1000
Industrial area	10, 100, 1000
Highway and railway	10, 100, 1000
Main river and canal	5, 50, 500
Grassland and arable land	5, 50, 500
Park and recreation zone	5, 50, 500
Military zone	5, 50, 500
Forest	1
Marshland	1
Heathland	1

For the RDA, we first calculated Nei genetic distances (Nei 1972) among populations in GENALEX v.6.4.1 and performed principal coordinate analyses (PCoA) on these distances using PrCoord 1.0 in the Canoco 4.5 software package (ter Braak and Šmilauer 2002). All resulting PCoA axes with positive eigenvalues were used as the dependent variables in the RDA. The purely spatial effect was investigated using distance-based Moran Eigenvector Maps (db-MEM) (Dray et al. 2006; Legendre and Legendre 2012). Also for the least cost models, MEMs were calculated based on pairwise least-cost distances between all sampled locations. Since we had no detailed landscape information for the Elsenborn region, this population was omitted from these analyses. For comparative purposes, the other RDA analyses were performed with and without this population. For the road infrastructure models, the position of sampled locations relative to the linear infrastructure was coded as a binary (dummy) explanatory variable. A population is thus located on one side of the infrastructure (having value=1) or on the other side (value=0). Having only sampled ten populations, the number of highways we can test is rather limited, as we need a minimum of two, but preferably more, populations on each side of the investigated infrastructure. We chose the E34 highway and the E313 plus E314 highways (with the E313 coinciding completely with the large and broad Albert canal), as they represent the most obvious testable linear barriers (Fig. 1). We chose not to include Mantel tests, as these are generally weak at detecting spatial patterns (Legendre et al. 2005; Gilbert and Bennett 2010; Legendre and Fortin 2010).

We used the double stop criterion of Blanchet et al. (2008) for RDA model selection, which consists of testing the significance of the global model (including all explanatory variables), and only continuing to forward selection of variables if this model is significant. Significance testing was done with 999 permutations. For the global RDA model, we report adjusted regression coefficients ( $R^2_{adj}$ ), thereby correcting for model overfitting (Peres-Neto et al. 2006). All RDA analyses were performed using the `rda` and `varpart` functions of the `vegan` package (Oksanen et al. 2011). In addition, the `forward.sel` function of the `packfor` package (Dray et al. 2006) was used. MEM spatial eigenfunctions were computed using the `create.MEM.model` function in the `PCNM` package (Legendre et al. 2010). All analyses were done in R v.2.15.1 ([www.r-project.org](http://www.r-project.org)).

## Results

### Genetic diversity and population structure

Altogether, 536 individuals, divided over the ten sampling regions, were genotyped. In total, 81 alleles were scored with the number of alleles per locus ranging from 3 (IDVGA2) to

**Table 2** Basic genetic parameters calculated for the different populations

Pop	Abbr.	N	NA	AR	PA	$F_{is}$	$H_e$	$H_o$
De Vart	DV	55	4.9	6.0	0	-0.005	0.604	0.607
Meerdaal	MD	55	5.5	6.97	1	0.074	0.595	0.551**
Molenbeersel	MB	47	5.0	6.50	0	0.055	0.597	0.564**
Balen	B	43	5.4	7.0	0	-0.008	0.630	0.635
Walenbos	WB	43	4.8	8.0	1	0.084	0.626	0.574*
Elsenborn	EB	58	6.5	9.13	6	0.005	0.659	0.656
Zwarte Beek	ZB	52	5.5	6.74	1	0.033	0.610	0.590
E19N	EN	70	5.0	5.80	0	0.031	0.655	0.635
Schijnvallei	SV	59	5.2	7.22	0	0.077	0.587	0.542**

*N* sample size, *NA* mean number of alleles per locus, *AR* allelic richness, *PA* private alleles, *F<sub>is</sub>* inbreeding coefficient, *H<sub>e</sub>* expected and *H<sub>o</sub>* observed heterozygosity with indication of a significant deviation from Hardy-Weinberg equilibrium at \**p*<0.05 or \*\**p*<0.001

17 (OarFCB304). Table 2 shows a summary of the main basic genetic parameters per population. The overall level of genetic diversity, measured as mean number of alleles per locus and allelic richness, was slightly higher in the population of Elsenborn than in the Flemish populations. However, except for Molenbeersel (Wilcoxon test,  $0.02 < p < 0.05$ ), the differences in allelic richness were not significant ( $p > 0.05$ ). Also, heterozygosity ( $H_e$  and  $H_o$ ) levels did not significantly differ ( $p > 0.05$ ) between populations, except for Molenbeersel and Elsenborn ( $0.02 < p < 0.05$ ). All populations except Bosbeekvallei, De Vart and Balen displayed a deficit of heterozygotes ( $F_{is} > 0$ ), and four populations (Schijnvallei, Meerdaal, Molenbeersel and Walenbos) showed a significant deviation from HWE at several of the loci. The overall  $F_{st}$  on the whole data set was 0.075, while pairwise  $F_{st}$ -values (Table 3) ranged from 0.031 (E19N with Balen) to 0.14 (Schijnvallei and De Vart with Elsenborn). All  $F_{st}$  values were significant at the  $p < 0.001$  level. From the 45 pairwise comparisons, 16 (36 %) lay between 0.08 and 0.14 indicating moderate differentiation. Pairwise  $D_{est}$ -values were largely comparable with the  $F_{st}$  values (Table 3) but here, 56 % of the pairwise values were higher than 0.08.

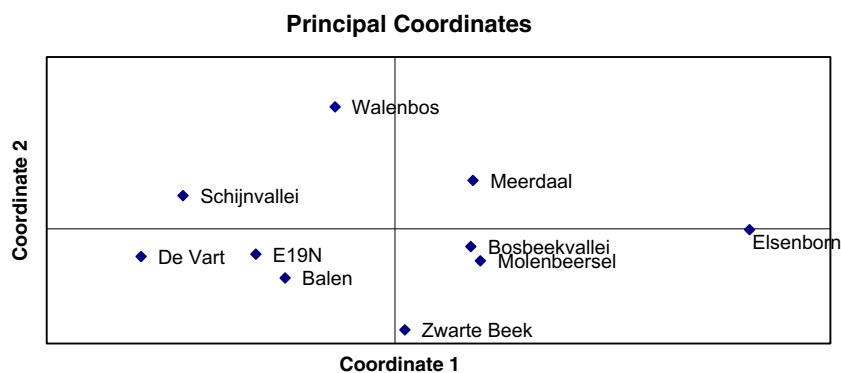
A principal coordinate analysis at the level of the individuals did not show a clear structure (Online Resource 3). At the population level, a pattern of gradual transition from the south-eastern, relict populations (Elsenborn, Molenbeersel and Meerdaal) towards the north-western, more recent ones (E19N, Schijnvallei and Balen), can be distinguished (Fig. 2), with the first principal coordinate largely separating both groups of populations. The same pattern is also reflected in the pairwise  $F_{st}$ -values between these populations (Table 3). The mean  $F_{st}$ -values amongst the relict populations are all lower than the values of relict populations compared to the younger ones. This pattern is also further supported by Bayesian analysis using STRUCTURE. The most likely partitioning of the genetic diversity based on delta K was obtained with  $K=2$ , while a weaker structure was obtained with  $K=4$  (Table 4 and Online Resource 1 and 4). For  $K=2$ , the majority of the individuals of Elsenborn, Meerdaal, Bosbeekvallei and Molenbeersel belong to one cluster, while De Vart, Balen, Walenbos, E19N and Schijnvallei form a second cluster. For  $K=4$ , more than 95 % of the deer from Elsenborn belong to one distinct cluster. A second cluster comprises most individuals from Meerdaal (89 %) together with around one third of the animals from the nearby

**Table 3** Pairwise  $F_{st}$  (below diagonal) and  $D_{est}$  (above diagonal) values between populations

	BBV	DV	MD	MB	B	WB	EB	ZB	EN	SV
Bosbeekvallei	–	0.106	0.065	0.063	0.055	0.080	0.090	0.041	0.104	0.111
De Vart	0.110	–	0.161	0.113	0.027	0.050	0.218	0.059	0.033	0.051
Meerdaal	0.068	0.107	–	0.057	0.120	0.083	0.097	0.077	0.128	0.105
Molenbeersel	0.056	0.097	0.059	–	0.054	0.119	0.084	0.044	0.098	0.098
Balen	0.061	0.028	0.085	0.060	–	0.065	0.163	0.035	0.039	0.077
Walenbos	0.083	0.075	0.072	0.078	0.060	–	0.140	0.116	0.063	0.082
Elsenborn	0.073	0.139	0.076	0.057	0.100	0.097	–	0.091	0.175	0.211
Zwarte Beek	0.059	0.061	0.060	0.046	0.037	0.095	0.074	–	0.069	0.091
E19N	0.070	0.040	0.089	0.067	0.031	0.061	0.102	0.056	–	0.042
Schijnvallei	0.088	0.044	0.084	0.099	0.061	0.072	0.140	0.082	0.043	–

Abbreviations are as in Table 2

**Fig. 2** Principal coordinate analysis of the populations. The first and second coordinate account respectively for 60.15 and 16.16 % of the total variation



Walenbos population. The other populations are divided over two additional clusters largely reflecting the geographic origins with all western populations in one cluster and the more eastern populations in another cluster. Only Balen and Walenbos have a substantial number of individuals in two clusters. According to the assignment probabilities computed in GENECLASS, 70 % of all individuals could be assigned to a specific population with a probability >0.7. Of these animals, between 70 and 98 % were correctly assigned to the predefined population (Table 5) except for Balen (59 %). When considering an assignment probability of more than 90 %, the percentage of correctly assigned individuals dropped below 20 % for three of the populations (De Vart, Balen and Zwarte Beek) and only stayed above 70 % for Meerdaal (73 %) and Elsenborn (83 %) (data not shown). In total, ten animals were excluded from their population of origin, two in Meerdaal and Molenbeersel and one each in the other populations except for Balen en Walenbos where there were no exclusions. Three individuals (one from Molenbeersel (exclusion threshold 0.001) and the ones from Elsenborn (exclusion threshold 0.008) and Schijnvallei (exclusion threshold 0.004)) were excluded from all populations.

#### Spatial population structure

The redundancy analysis using purely spatial distance-based Moran eigenvector maps, and including Elsenborn, yielded two significant db-MEM-variables in the global model which together explained 18.2 % of the genetic data ( $p=0.023$ ) (Table 6). Both variables were retained after forward selection. When Elsenborn was excluded, four db-MEM-variables were modelled ( $R^2_{adj}=0.302, p=0.011$ ). Of these four variables, the first two were sufficient to significantly ( $p=0.001$ ) explain the adjusted variance. A separate redundancy analysis considering only the highways E34 and E313/E314 (plus Albert canal) (Fig. 1) as explanatory variables significantly explained 23.4 % (including Elsenborn) and 28 % (excluding Elsenborn) of the genetic variance (Table 6). None of the three least cost models (LCM10, LCM100 and LCM1000), having respectively five, three and five MEM-variables in the global model, yielded a significant overall result.

#### Discussion

Our results show that the roe deer populations sampled in Flanders display genetic diversity parameters that are slightly lower compared to the stable and large population from Elsenborn. Mean number of alleles, allelic richness and number of private alleles were higher in the Walloon population and comparable to other studies on roe deer using nearly the same set of markers (Coulon et al. 2004; Thulin 2006). The obtained genetic diversity parameters are also in the same range as those obtained for other ungulate species such as red deer in France (Dellicour et al. 2011), Belgium (Frantz et al. 2012), Germany (Kuehn et al. 2003) and Denmark (Nielsen et al. 2008) or moose in Norway (Haanes et al. 2011). Other estimates such as observed and expected heterozygosity were similar among all populations studied, except for Molenbeersel, which significantly differed from Elsenborn. Four Flemish populations (Schijnvallei, Meerdaal, Molenbeersel and Walenbos) nevertheless showed a significant heterozygosity deficit and deviation from Hardy-Weinberg equilibrium. However, when excluding the juveniles from the data set, Meerdaal no longer showed a deviation from HWE (data not shown). The deficit of heterozygotes

**Table 4** Assignment of roe deer individuals to genetic clusters using STRUCTURE

Population	cluster 1	cluster 2	cluster 3	cluster 4
Bosbeekvallei	0.008	0.032	<i>0.917</i>	0.043
De Vart	0.003	0.006	0.023	<i>0.968</i>
Meerdaal	0.004	<i>0.887</i>	0.015	0.093
Molenbeersel	0.188	0.004	<i>0.615</i>	0.193
Balen	0.007	0.011	<i>0.456</i>	<i>0.526</i>
Walenbos	0.009	<i>0.351</i>	0.015	<i>0.626</i>
Elsenborn	<i>0.958</i>	0.007	0.027	0.007
Zwarte Beek	0.024	0.010	<i>0.794</i>	0.172
E19N	0.003	0.007	0.017	<i>0.973</i>
Schijnvallei	0.004	0.070	0.018	<i>0.908</i>

Assignment values higher than 30 % are indicated in *italic*



**Table 5** Assignment of roe deer individuals to predefined populations using GENECLASS

	BBV	DV	MD	MB	B	WB	EB	ZB	EN	SV
Bosbeekvallei	<i>89.7</i>	0	2.6	2.6	2.6	0	0	2.6	0	0
De Vart	0	<i>90.6</i>	0	0	0	3.1	0	3.1	0	3.1
Meerdaal	2.1	0	<i>91.5</i>	0	0	2.1	0	2.1	2.1	0
Molenbeersel	3	0	0	<i>90.9</i>	0	0	0	6	0	0
Balen	9.1	9.1	0	9.1	59.1	4.5	0	9.1	0	0
Walenbos	0	0	0	3.6	0	<i>92.9</i>	0	0	0	3.6
Elsenborn	0	0	0	0	0	0	<i>98</i>	0	2	0
Zwarte Beek	6.1	0	3	6.1	3	0	3	<i>69.7</i>	6.1	3
E19N	0	6.7	0	0	0	0	0	0	<i>85.7</i>	6.7
Schijnvallei	0	13.5	2.7	0	0	8.1	0	0	5.4	<i>70.3</i>

Correct assignments of 70 % and higher are indicated in *italic*  
Abbreviations are as in Table 2

may be either due to population substructure or inbreeding in a panmictic population. Within Flanders, there were also no significant differences in overall genetic diversity between the known relict populations (Molenbeersel and Meerdaal) and more recently established populations (E19N, Schijnvallei and Balen). The fact that we found no evidence for lower genetic diversity in more westwards, recently recolonized regions than in known relict populations, indicates that recolonization involved sufficient founding sizes, or that gene flow has already replenished the genetic diversity. Unreported translocations or release of captive animals may also have contributed.

Several of the roe deer populations examined show pairwise differentiation values above 0.08, indicating moderate to quite strong differentiation. The highest values are found between the relict and recent populations. This is reflected in a PCoA plot showing a smooth transition from south-east to north-west. Assignment tests using STRUCTURE and GeneClass also support the hypothesis of genetic differentiation and population structure in Flanders, although the power to uncover very recent genetic processes (<10 generations) with only ten genetic markers is rather low. STRUCTURE recognizes four main genetic clusters. From the perspective of ancient occurrence and current geographic distribution of roe deer in Flanders, a population structure with four clusters makes sense and might well correspond with the presence of remnant populations in the past and the recent recolonization, showing that expansion did not occur over large distances. This is in agreement with the general knowledge that roe deer do not disperse over large distances (Linnell et al. 1998), and that barriers may interrupt migration and gene flow (Coulon et al. 2006; Kuehn et al. 2007). The existence of a subtle population structure with little migration is further confirmed by the assignment tests with GeneClass which correctly reassigned a large part of the individual animals to the predefined populations except for Balen. This is not surprising as Balen is located in the middle of the study area and is more connected by suitable landscape with other adjacent

sampled populations such as Zwarte Beek. As a result, gene flow among these centrally located populations might be higher. Three animals could not be assigned to any of the populations examined. It remains unclear from where they originate and whether they are natural migrants, especially since we cannot exclude the possibility of undocumented translocations.

Overall, our results point towards the existence of a quite clear local population structure at small spatial scale and rather restricted gene flow among the sampled units. To test whether this could be correlated with spatial distance, physical barriers or landscape features, we performed redundancy analyses. The RDA showed a significant and relatively strong effect of purely spatial structure (distance) that likely reflects spatial autocorrelation: nearby units are more similar than distant units. On the other hand, the position of populations relative to broad and major transportation infrastructure (two highways and a canal) had nearly as much explanatory power, while this model (consisting of just two binary variables) was much simpler. In contrast, the rather complicated least cost

**Table 6** Results of the redundancy analyses using different explanatory models with (10 populations) or without (9 populations) Elsenborn

Expl. Var.	Pops	Var.	R <sup>2</sup> <sub>adj</sub>	Sel. Var.	R <sup>2</sup> <sub>adj</sub>
Space	9	4	0.302 (0.011)	2	0.302 (0.001)
	10	2	0.182 (0.023)	2	0.182 (0.023)
Roads/canals	9	2	0.280 (0.023)	2	0.280 (0.023)
	10	2	0.234 (0.010)	2	0.234 (0.010)
LCM10	9	5	NS (0.917)		
LCM100	9	3	NS (0.511)		
LCM1000	9	5	NS (0.990)		

*Expl. Var.*: explanatory variant used in the model; *Pops*: number of populations; *Var.*: total number of variables in the global model; *Sel. Var.*: number of selected variables; *R<sup>2</sup><sub>adj</sub>*: adjusted explained variance with the *p* value between brackets. Negative values are set to zero and are denoted as NS (not significant)

models had no explanatory power. Possibly, the fine-grained landscape structure, with a complex mosaic of both suitable and unsuitable patches, has no strong effect in our study approach. It is probable that such models are more appropriate for spatially-explicit, individual-based sampling designs at smaller spatial scales. Overall, our results suggest that the considered transportation infrastructure truly hinders gene flow. Individual-based studies have shown this elsewhere too for roe deer (Kuehn et al. 2007; Hepenstrick et al. 2012) or other large mammals (Perez-Espona et al. 2008; Frantz et al. 2012). Also, studies using telemetry showed that in the case of red deer, highways form a barrier that is hardly crossed (Licoppe 2006). In our study, the potential impact of anthropogenic barriers is further exemplified by the populations of Walenbos and Meerdaal, which despite their physical proximity, are relatively strongly differentiated from each other ( $F_{st}=0.072$ ,  $D_{est}=0.083$ ) and only partially belong to the same genetic cluster. Although it could not explicitly be tested, this may be due to the presence of the E40 highway and an accompanying railway separating the two populations from each other. In contrast, Balen and Zwarte Beek which are comparably close together, only show low differentiation levels ( $F_{st}=0.037$ ,  $D_{est}=0.035$ ), most probably because there are no harsh physical barriers separating them and there is sufficient suitable habitat in the landscape matrix between them. In this respect, it is unfortunate that we don't have individual coordinates of the animals' analyses. An individual-based approach would make it much more accurate to assess the impact of barriers at the one hand and for example green highway overpasses or corridors at the other hand on gene flow, especially in such a highly fragmented landscape as Flanders.

From this study, we cannot conclude that the twentieth century decline of roe deer, and its subsequent recovery and western recolonization, were accompanied by a noticeable reduction in genetic diversity. The spatially structured genetic diversity likely reflects the combined effect of spatial autocorrelation (isolation-by-distance) and major transportation infrastructure that reduces gene flow. As such, barriers and other landscape features (such as woodland) that influence dispersal have a genetic and biological significance that should be taken into account when defining management units. Rather than purely administrative boundaries, also genetic and other biological components should be taken into account for the delimitation of biological meaningful management units.

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